

Probabilistic Coreference in Information Extraction

Andrew Kehler
SRI International
333 Ravenswood Avenue
Menlo Park, CA 94025
kehler@ai.sri.com

Abstract

Certain applications require that the output of an information extraction system be probabilistic, so that a downstream system can reliably *fuse* the output with possibly contradictory information from other sources. In this paper we consider the problem of assigning a probability distribution to alternative sets of coreference relationships among entity descriptions. We present the results of initial experiments with several approaches to estimating such distributions in an application using SRI's FASTUS information extraction system.

1 Introduction

Natural language information extraction (IE) systems take texts containing natural language as input and produce database templates populated with information that is relevant to a particular application. These records may be fed as input to a downstream system for which the IE system is only one of several sources of information. In such a scenario, the downstream system must *fuse* the incoming information from each of its sources, requiring the resolution of conflicts. To accomplish this, the fusion system must know the reliability of the information received from each source; in this way unreliable information from one source can be disregarded in favor of highly reliable information from another.

Figure 1 exhibits this scenario with a typical IE system such as SRI's FASTUS system (Hobbs et al., 1996). The IE system has two components. The first component consists of a series of phases that recognize domain-relevant patterns in the text and create templates representing event and entity descriptions from them. The second component merges templates created from different phrases in the text that overlap in reference. The resulting set of templates

constitutes a formal description of the state of affairs as described in the text with respect to the application specification, which is then fed to the downstream system.

As part of determining this state of affairs, the IE system must create templates describing the relevant entities that are reported on. This requires determining when two or more templates describe the same entity, as templates created from corefering phrases need to be merged. We have performed an informal study of FASTUS's processing of a set of texts which indicates that the merging phase is where most of the ambiguities (as well as most of the errors) lie. However, most IE systems, including FASTUS, have pursued a deterministic strategy for merging and report only a single possible state of affairs. This limitation makes it difficult for a downstream system to fuse the information with possibly contradictory information from other sources, as no information about the IE system's certainty of the results is passed along, nor is information about possible alternative states of affairs and their associated levels of certainty.

In this paper, we consider the problem of assigning a probability distribution to alternative sets of coreference relationships among entity descriptions. We present the results of initial experiments with several approaches to estimating such distributions in an application using FASTUS.

2 Overview of the Problem

Let us consider an example text of the sort that we encounter in our application:¹

¹The texts in our application are messages consisting of free text, possibly interspersed with formatted tables or charts which themselves may contain natural language fragments that require analysis. While this example is shorter than most texts in our corpus, the relevant free text portions of the messages are typically no longer than a few paragraphs. The style displayed in this example is fairly typical, although in some cases the sentence struc-

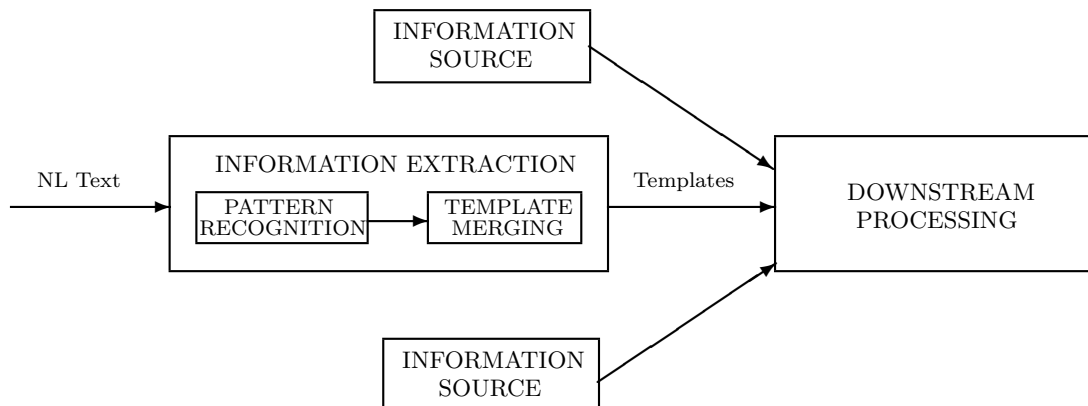


Figure 1: A Scenario Employing an Information Extraction System

Subj: *Kinston Military Rail Depot*

A rail depot was found 100 km southwest of the capitol of Raleigh, consisting of extensive admin and support areas (similar to the ammunition depot in Fairview), two material storage areas, extensive transshipment facilities (some of which are under construction immediately east of the depot), and several training areas.

We focus on the four mentions of depots in the text, which are highlighted with italics. The pattern matching phases of FASTUS produce templates similar to those shown in Figure 2.

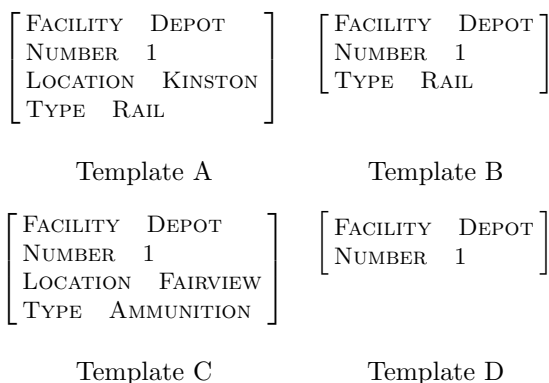


Figure 2: Templates Representing Depots Mentioned

We will refer to a set of templates that have potential coreference relationships among them as a cell is more telegraphic.

coreference set,² and possible partitions of coreferential templates in the set as *coreference configurations*. In the coreference set containing templates A, B, C, and D, system knowledge external to the probabilistic model indicates that the type *Ammunition* in template C is not compatible with the type *Rail* in A and B; therefore these are taken *a priori* to be non-coreferential. Given these incompatibilities, seven possible coreference configurations remain. Template names grouped within parentheses are taken to be mutually coreferring; we will refer to such a grouping as a *cell* of the coreference configuration.

1. (A B D) (C)
2. (A B) (C D)
3. (A B) (C) (D)
4. (A D) (B) (C)
5. (B D) (A) (C)
6. (C D) (A) (B)
7. (A) (B) (C) (D)

The first of these configurations expresses the correct coreference relationships for the example.

Given a coreference set of templates, possibly coupled with a list of template pairs known a priori not to corefer, the task is to assign a probability distribution over the possible coreference configurations for that set.

Relationship to Past Work While there have been previous investigations of empirical approaches to coreference, these have generally centered on the task of assigning correct referents for *anaphoric* expressions (Connolly, Burger, and Day, 1994; Aone and Bennett, 1995; Lappin and Leass, 1994; Dagan

²Templates A, B, C, and D constitute the only coreference set in this example, since none of the other NPs (e.g., the various “areas” mentioned) are compatible with any of the others. In general, however, a text can give rise to any number of distinct coreference sets, each of which will be assigned its own probability distribution.

and Itai, 1990; Dagan et al., 1995; Kennedy and Boguraev, 1996a; Kennedy and Boguraev, 1996b). The current task deviates from that problem in several respects. First, in our task, all coreference relationships among templates are modeled regardless of the “referentiality” of the phrases that led to their creation. For instance, indefinites will sometimes corefer with a previously described entity; a typical case is illustrated by the coreference between the indefinite “a rail depot” and the depot introduced in the subject line in the example passage. Also, entities described with bare plurals are commonly found to be coreferential with other entities, in addition to cases in which they have their more standard generic meanings. On the other hand, definite noun phrases are often not referential to items evoked in the text (e.g., “the ammunition depot in Fairview”). Determining when such expressions are discourse-anaphoric is part of the task; this information is generally not known to the system a priori.

Second, the results of this task will be evaluated by the probability assigned to the correct state of affairs with respect to an entire coreference set, and not by the number of correct antecedents assigned to anaphoric expressions. Modeling at the level of coreference sets ensures that the probabilities are consistent when considering the global state of affairs being described in the text. Furthermore, the role of probabilities for this application goes beyond selecting the correct coreference relationships – the probability assigned to an alternative will be central in determining how the downstream system will weigh it against information from other sources during data fusion. A system that assigns a probability of 0.9 to correct answers is more successful than one that assigns a probability of 0.6 to them.

The Limitations of IE Systems The properties of typical IE systems such as FASTUS also make this task challenging. For one, successful modeling of coreference relationships is hampered by the crudeness of the representations used. The templates that are created are fairly shallow and may be incomplete. A reliance on detailed information about the context can prove detrimental if such information is often missed by the system. Also, FASTUS also does not build up complex representations for the syntax and semantics of sentences, placing limits on the extent to which such information can be utilized in determining coreference. Lastly, there are the inaccuracies that result from processing real text. The pattern matching phases of FASTUS may intermittently misanalyze phrases that serve as antecedents for subsequent referring expressions. Therefore, for

example, with respect to an identified coreference set, it may be correct to place a referential pronoun in its own cell (implying that it does not corefer with anything), simply because system error caused its antecedent not to be included in the set.

Outline of the Approach The number of coreference configurations over which a distribution is to be assigned depends on the number of templates in the coreference set, and the set of a priori constraints against coreference between some of its members. As there are many scenarios that will never be encountered in a corpus of training data of any reasonable size, it would be hopeless to attempt to estimate a conditional distribution for each possibility directly. To make matters worse, training data comes at a cost, as keys have to be coded by hand. One of the goals of this effort is to allow the ability to train up probabilities in new domains quickly, which requires an approach that is successful with a limited amount of training data.

However, it would be reasonable to expect that we have enough data to estimate distributions for coreference sets with only two members. This suggests a two-step approach. First, we develop a general model of coreference between any two templates, and apply it to pairwise combinations of templates in a given coreference set without regard to the other templates in the set. We then utilize a method for combining the resulting probabilities to form a distribution over all the possible coreference configurations. We describe our method for modeling probabilities between pairs of templates in the next section, and describe two methods for deriving a distribution over the coreference configurations in Section 4. We report on an evaluation and comparison of the approaches in Section 5.

3 Training A Model for Pairs of Templates

Our first task is to derive a model for determining the probability that two templates corefer, conditioned on various characteristics of the context. For this we employ an approach to maximum entropy modeling described by Berger et al. (1996).

Maximum Entropy Modeling Suppose we wish to model some random process, such as that which determines coreference between two templates generated by an IE system, based on various characteristics of the context that influence this process, such as the content of the templates themselves, the form of the natural language expressions from which the templates were created, and the distance between

those expressions in the text. We refer to the collection of such characteristics for a given example as its context x , and the value denoting the output of the process as y . We can define a set of binary *features* that relate a possible value of a characteristic of x with a possible outcome y , i.e., whether the two templates corefer ($y = 1$) or not ($y = 0$). For example, a feature $f_1(x, y)$ pairing the characteristic of S and T having identical slot values with the outcome that they corefer would be defined as follows.

Binary Feature $f_1(x, y)$:

$$f_1(x, y) = \begin{cases} 1 & \text{if S and T have identical} \\ & \text{slot values and S and T corefer} \\ 0 & \text{otherwise} \end{cases}$$

From these features we can define *constraints* on the probabilistic model that is learned, in which we assume that the expected value of the feature with respect to the distribution of the training data (p_d) holds with respect to the general model (p_m).

Constraints:

$$\sum_{x, y} p_d(x, y) f(x, y) = \sum_{x, y} p_d(x) p_m(y|x) f(x, y)$$

Given that we have chosen a set of such constraints to impose on our model, we wish to identify that model which has the maximum entropy – this is the model that assumes the least information beyond those constraints. Berger et al. (1996) show that this model is a member of an exponential family with one parameter for each constraint, specifically a model of the form

$$p(y|x) = \frac{1}{Z(x)} e^{\sum_i \lambda_i f_i(x, y)}$$

in which

$$Z(x) = \sum_y e^{\sum_i \lambda_i f_i(x, y)}$$

The parameters $\lambda_1, \dots, \lambda_n$ are Lagrange multipliers that impose the constraints corresponding to the chosen features f_1, \dots, f_n . The term $Z(x)$ normalizes the probabilities by summing over all possible outcomes y . Berger et al. (1996) demonstrate that the optimal values for the λ_i 's can be obtained by maximizing the likelihood of the training data with respect to the model, which can be performed using their *improved iterative scaling* algorithm.

In practice, we will not want to incorporate constraints for all of the features that we might define,

but only those that are most relevant and informative. Therefore, we use a procedure for selecting which of our pool of features should be made *active*. At each iteration, the algorithm approximates the gain in the model's predictiveness that would result from imposing the constraints corresponding to each of the existing inactive features, and selects the one with the highest anticipated payoff. Upon making this feature active, the λ_i 's for all active features are (re)trained so that the constraints are all met simultaneously. The feature selection process is iterated until the approximate gain for all the remaining inactive features is negligible.

Characteristics of Context for Template

Coreference We now need a set of possible characteristics of context on which the algorithm could choose to conditionalize in deriving the probabilistic model. For our initial experiments, we utilized a set of easily computable, but fairly crude, characteristics.³ These characteristics fall into three categories. In what follows, we take S and T to be arbitrary templates where the natural language expression from which T was created appears later in the text than the expression from which S was created.

The first category relates to the contents of the templates themselves. We model the relationship between S and T as one of the following: S and T have *identical slot values*, S is *properly subsumed* by T, S *properly subsumes* T, or S and T are *otherwise consistent*. For instance, in our example in Section 2, template A is properly subsumed by template B, and A, B, and C are all properly subsumed by D, since in each case the latter template is more general than the former. We also have a binary characteristic for S and T having at least two (non-nil) slot values in common. Finally, we have a characteristic for modeling when the values of the NAME slot of a template are both multi-worded and identical; this is a crude heuristic for identifying matching unique identifiers.

The second category of characteristics relates to the form of reference used in the expression from which T was created, specifically whether it was de-

³One could imagine a variety of more detailed and informative characteristics of context than those used here. However, in performing these experiments, we are interested in how far we can get with a fairly simple strategy that will port relatively easily to new domains, rather than relying heavily on information that is specific to our current domain. A fairly coarse-grained set of characteristics also allows us to restrict ourselves to a relatively small set of training data; likewise we will not want to encode a large set of data for each new domain.

Template S	Template T	Probability
A	B	0.671
A	D	0.505
B	D	0.752
C	D	0.504

Table 1: Pairwise Probabilities for Example Coreference Set

scribed with an *indefinite* phrase, a *definite* phrase (including pronouns), or *neither* of these (e.g., a bare, non-pronominal noun phrase). In the case of definite expressions, we also consider the recommendations of a distinct coreference module within FAS-TUS. We have a characteristic representing whether the potential antecedent is the *preferred* antecedent,⁴ a *non-preferred, but possible* antecedent, or not on the list of possible antecedents.⁵

The final category of characteristics relates to the distance in the text between the expressions from which S and T were created, which we categorize as being in one of five equivalence classes: *very close*, *close*, *mid-distance*, *far away*, and *very far away*. These distances are measured crudely (i.e., by character length) so as not to be dependent on the accuracy of methods for identifying more complex boundaries (e.g., clause, sentence, and discourse segment boundaries).

The results of training the maximum entropy models are discussed in Section 5. To illustrate the approaches described in the next section, we will use the probabilities for the templates from the example passage in Section 2, shown in Table 1, which were produced from the parameters induced from one of the training sets.

4 Inferring a Model for Coreference Sets

We now have a method for obtaining a model that assigns probabilities to the pairs of templates (henceforth, “pairwise probabilities”) in a coreference set that can possibly corefer. If there are only two templates in the coreference set, then we have the distri-

⁴Preferred reference is a transitive relation, that is, template S is treated as a preferred referent of template T if there is a chain of preferred referents linking them, e.g., if there is a template R that is the preferred referent of T and template S is the preferred referent of R.

⁵Although we do not model information about the surface positions of the expressions from which S and T were created within their respective sentences, the coreference module does take such information into account in determining likely antecedents of definite expressions.

bution we seek. However, if there are more than two templates, we must utilize the pairwise probabilities to derive a distribution over the members of the set of coreference configurations. In the following sections, we describe two approaches to recovering such a distribution, followed by a description of two baseline metrics. An evaluation of these approaches is then given in Section 5.

4.1 An Evidential Reasoning Approach

The first approach we describe uses the pairwise probabilities as sources of evidence that inform the choice of model for the coreference sets. The list of coreference configurations for our example passage are repeated below; we will refer to these configurations by their corresponding numbers.

1. (A B D) (C)
2. (A B) (C D)
3. (A B) (C) (D)
4. (A D) (B) (C)
5. (B D) (A) (C)
6. (C D) (A) (B)
7. (A) (B) (C) (D)

We recast a probability that two templates S and T corefer as a *mass distribution* over two members of the power set of coreference configurations, namely the set containing exactly those configurations in which S and T occupy the same cell, and the set containing those in which they do not. For instance, the probability that A and B corefer was determined to be 0.671; mapping this to corresponding sets of coreference configurations results in the mass distribution m_{AB} in which

$$m_{AB}(\{\text{Configs } 1, 2, 3\}) = 0.671$$

and

$$m_{AB}(\{\text{Configs } 4, 5, 6, 7\}) = 0.329$$

This mass distribution can be seen as representing the beliefs of an observer who only has access to templates A and B, and who is therefore ignorant about their relationship to C and D. We can view the other pairwise probabilities for the coreference set in the same manner.

In the best of all worlds, we might identify a model that is consistent with the mass distributions provided by all the pairwise probabilities. However, such a model may not, and often will not, exist. This is the case for the pairwise probabilities in our example, which can be seen most easily by considering only templates A, C, and D. The probability of A and D corefering is 0.505 and of C and D corefering is 0.504. Because we know that A and C cannot corefer, the coreference configurations in which A and D corefer and the configurations in which C

and D corefer are mutually exclusive. Therefore, there would have to be a distribution that assigns 0.505 of probability mass to a set of configurations that is mutually exclusive from a set that is assigned 0.504 of probability mass. Obviously, this cannot be done with a set of probabilities that add up to 1.

This inconsistency arises from the manner in which the pairwise probabilities are estimated. The probability of coreference between templates situated similarly to A and D may be 0.505 with respect to all contexts in the training data, however it is almost certainly not this high with respect to the subset of cases in which a template similar to C is similarly situated. The same reasoning applies to the probability of C and D coreferring in light of the existence of A. Unfortunately, the existence of templates other than the pair being modeled is the type of conditional information for which we have little hope of accounting in a general and statistically significant manner.

Therefore, we may be left with a series of mass distributions defined over sets of coreference configurations that are in inherent conflict. Instead of viewing these distributions as constraints on the underlying probabilistic model, we view them as sources of evidence. The question is then how to take these sources into account, given that they may be partially contradictory. *Dempster's Rule of Combination* (Dempster, 1968) provides a mechanism for doing this. Dempster's rule combines two mass distributions m^1 and m^2 to form a third distribution m^3 that represents the consensus of the original two distributions; the new mass distribution in effect leans toward the areas of agreement between the original distributions and away from points of conflict. Dempster's rule is defined as follows:

$$m^3(A_k) = \frac{1}{1 - \kappa} \sum_{A_i \cap A_j = A_k} m^1(A_i)m^2(A_j)$$

in which

$$\kappa = \sum_{A_i \cap A_j = \emptyset} m^1(A_i)m^2(A_j)$$

The A_i in our case are members of the power set of possible coreference configurations. In our example above, m_{AB} assigns probability mass to two such A_m , the set containing configurations 1, 2, and 3, and the set containing configurations 4, 5, 6, and 7.

The value κ is called the *conflict* between the mass distributions being combined; it provides a measure of the degree of disagreement between them. When $\kappa = 0$, the original distributions are compatible; when $\kappa = 1$, they are in complete conflict and the

result is undefined. When $0 < \kappa < 1$, some conflict between the distributions exists; Dempster's rule has the effect of focusing on the agreement between the distributions by eliminating the conflicting portions and normalizing what remains.

We can therefore use Dempster's Rule to resolve the conflict between the pairwise probability distributions to generate a distribution over the coreference configurations. Because we have pairwise probabilities for each possibly coreferring pair in the coreference set, it turns out that the Dempster solution is more easily stated and computed here than in the general case. The solution is identical to the one that results when the probabilities of all the relevant pairwise relations (indicating either coreference or not) are multiplied, normalized by the amount of probability mass assigned to coreference configurations that are impossible because coreference is transitive. For instance, the probability for the coreference configuration ((A B) (C)) is initially computed to be⁶

$$p(A =_c B) * p(A \neq_c C) * p(B \neq_c C)$$

However, using this method, impossible combinations (e.g., $A =_c B, B =_c C, A \neq_c C$) will also receive positive probability mass. If we normalize the probabilities of possible combinations by distributing the sum of the probability assigned to all impossible combinations, the result is the same as that gotten by iteratively combining the pairwise distributions using Dempster's Rule.

The resulting distribution for our example is:

1. (A B D) (C) = .383
2. (A B) (C D) = .184
3. (A B) (C) (D) = .123
4. (A D) (B) (C) = .062
5. (B D) (A) (C) = .125
6. (C D) (A) (B) = .061
7. (A) (B) (C) (D) = .061

In motivating our approach, we noted that we cannot expect to have the amount of training data necessary to directly estimate distributions for all the possible scenarios with which we may be confronted. Limiting ourselves to modeling probabilities between pairs of templates, however, leads to inconsistencies because of the failure to take into account the crucial information provided by the existence of other compatible templates. Dempster's Rule can be seen as a very coarse-grained approach to conditioning on context in this regard. The contributions of the pairwise models are conditioned not on the existence of other

⁶We use the notation $=_c$ to indicate coreference.

templates in context, but by virtue of the existence of conflicting models derived from those templates. For instance, the pairwise probability of coreference between C and D was originally 0.504, which might be reasonable if those were the only two templates generated from the text.⁷ However, the probability that C and D corefer in the final distribution is only 0.245, the sum of the probabilities of the two partitions in which C and D occupy the same cell. This adjustment results from the existence of templates A and B: the fact that template D has a high probability of coreferring with each, combined with the fact that template C is incompatible with each, reduces the likelihood that C and D corefer. Therefore, the preferences for particular coreferential dependencies can change when considering the larger picture of possible coreference sets.

In practice, coreference sets that are significantly larger than the one we have considered here can lead to an explosive number of possible coreference configurations. We have implemented simple methods for pruning very low probability configurations during processing and for smoothing the resulting distribution. The latter step is accomplished, when necessary, by eliminating certain low-probability configurations at the end of processing. The probability mass from these configurations is distributed uniformly over all the possible configurations that have been eliminated. While this is unlikely to be the best strategy for smoothing from the standpoint of probabilistic modeling, we are constrained by the number of alternatives we can report to the downstream system. Smoothing in this way allows us to report only the coreference configurations with non-negligible probability, along with a single probability that is assigned uniformly to the remainder of the possible configurations.

4.2 A Model Based on Merging Decisions

The second approach we consider models the likelihood of correctness of decisions that a template merger such as the one used in FASTUS would make in processing a text. To illustrate, consider the case in our example in which the probability of the coreference configuration ((A B D) (C)) is determined. The merger would make the following decisions in deriving such a configuration, in which the notation “B&A” represents the template that results from

⁷Actually this number is lower than it would have been, because template B was identified as the preferred antecedent for template D instead of template C. If C and D were the only two templates generated, then C would have been identified as the preferred antecedent, thus raising the probability.

templates A and B having previously been merged.

1. B =_c A? → yes
2. C =_c B&A? → no
3. D =_c C? → no
4. D =_c B&A? → yes

We therefore model the probability of this coreference configuration as the product of each of the corresponding pairwise probabilities. Since we cannot model coreference involving objects that have resulted from previous (hypothetical) merges – the appropriate feature values for distance and form of referring expression would become unclear – we make the following approximation:

$$p(X =_c Y_1 \& \dots \& Y_n) \approx p(X =_c Y_n)$$

in which Y_n is the most recently created template in Y_1, \dots, Y_n .

Using the probabilities from Table 1,⁸ the probability assigned to ((A B D) (C)) would therefore be

$$p(B =_c A) * p(C \neq_c B) * p(D \neq_c C) * p(D =_c B) = 0.671 * 1 * (1 - 0.504) * 0.752 = 0.250$$

Note that unlike the evidential approach, the probability of the pair D and A coreferring does not come into play, given that coreference between D and B and between B and A has been factored in.

This approach yields a probabilistic model as given, that is, the probabilities sum to 1 without normalization. However, in certain circumstances the approximation above will generate probability mass for an impossible case, specifically when it is known a priori that X is incompatible with one of the templates Y_1, \dots, Y_{n-1} . For instance, if templates B and C in our example had been compatible (with A and C remaining incompatible), then the approximation above would assign positive probability mass to the coreference configuration ((A B C) (D)), because the zero probability of A coreferring with C would not come into play. Therefore we modify the above approximation to apply only if X and each of Y_1, \dots, Y_{n-1} are compatible; otherwise, the probability mass assigned is used for normalization. One can see that this can only improve the pure form of the model.

Using the pairwise probabilities from Table 1, the results of the model as applied to the example are:

⁸We use these probabilities for ease of comparison. In reality, the pairwise probabilities for this model were trained with an adapted set of training data as explained below, and so these numbers are in actuality a bit different.

1. (A B D) (C) = .250
2. (A B) (C D) = .338
3. (A B) (C) (D) = .083
4. (A D) (B) (C) = .020
5. (B D) (A) (C) = .123
6. (C D) (A) (B) = .166
7. (A) (B) (C) (D) = .020

4.3 Two Bases of Comparison

We compared the two learned models with two baseline models. First, as an absolute baseline, we compared the model with the uniform distribution, that is, the distribution that assigns equal probability to each alternative. We then sought a more challenging, yet straightforward baseline. We defined a simple, “greedy” approach to merging similar to the one used in FASTUS, in which merging of newly-created templates is attempted iteratively through the prior discourse, starting with the most recently produced object. Any unifications that succeed are performed. For instance, in the above example, the greedy method produces the configuration ((A B) (C D)), because A is compatible with B, C is not compatible with either, and D is compatible with C (with which merging would be attempted before the earlier-evoked templates B and A). Alternatively, in cases in which all of the templates in a coreference set are pairwise compatible, the greedy method will produce the configuration in which they are all coreferential.

We then calculated how often this approach yielded the correct results in each training set. We distinguished between three values: the percentage of correctness for coreference sets of cardinality 2 (call this p_2), the percentage for coreference sets of cardinality 3 (call this p_3), and the percentage for coreference sets of cardinality 4 or more (call this $p_{>3}$). The greedy model was defined such that the result of the greedy merging strategy is assigned the appropriate probability p_k , with the remainder of the probability mass $1 - p_k$ distributed uniformly among the remaining possible alternatives. (No alternatives were included that were *a priori* known to be impossible due to incompatibilities.)

For instance, in the first training set we describe below, $p_2=.571$, $p_3=.652$, and $p_{>3}=.344$ (the percentage for the whole training corpus was $p=.555$). If there are 4 templates, and 10 coreference configurations are possible, then the answer derived by the greedy strategy would receive probability .344, and the remaining 9 alternatives would receive probability $\frac{1-.344}{9} = .0729$. In the second training set we describe below, $p_2=.646$, $p_3=.600$, and $p_{>3}=.345$ (the percentage for the whole train-

ing corpus was $p=.549$), and in the third training set, $p_2=.628$, $p_3=.600$, and $p_{>3}=.280$ (the percentage for the whole training corpus was $p=.523$).

5 Experiments

5.1 Training the Maximum Entropy Models

For reasons described below, we trained separate pairwise probability models for each of the two approaches. We ran FASTUS over our development corpus, 72 texts of which produced coreference data. The texts gave rise to 132 coreference sets, and produced characteristics of context for 647 potential coreference relationships between pairs of templates. We created a key by analyzing the texts and entering the correct coreference relationships.

We created three splits of training and test data. In the first split, the training set contained 60 messages, giving rise to 110 coreference sets, and the test set contained 12 messages, giving rise to 22 coreference sets. In the second split, the training set contained 57 messages, giving rise to 102 coreference sets, and the test set contained 15 messages, giving rise to 30 coreference sets. The third test set was created by combining the first and second test sets. The training set contained 47 messages, giving rise to 88 coreference sets, and the test set contained 25 messages (the first two test sets overlapped by two messages), which gave rise to 44 coreference sets.

For training the maximum entropy model, only the sets of characteristics of context for pairwise coreference are relevant; the number of such sets differed between the two approaches as discussed below. The evaluations were performed on the test sets with respect to the final distribution generated for the coreference sets, with the result being measured in terms of the average cross-entropy between the model and the test data.

Data for the Evidential Model The evidential model utilizes the pairwise probabilities between all pairs of templates in a coreference set. Therefore, we used all such pairs in each training set to train the maximum entropy model. In the first training set, the 110 coreference sets gave rise to characteristics of context for 578 pairs of templates; in the second, the 102 coreference sets gave rise to characteristics for 581 pairs of templates. In the third training set, the 88 coreference sets gave rise to characteristics for 525 pairs of templates.

The maximum entropy algorithm selected similar sets of features to model in each case.⁹ Among the

⁹The following features represent the referenced char-

systems of λ_i values learned, negative values were learned for the features in which template S *properly subsumes* template T and in which S and T are *otherwise consistent*. These two features model the cases in which template T contains information not contained in template S, reflecting the fact that expressions referring to the same entity usually do not become more specific as the discourse proceeds. A positive value was learned for the feature modeling cases in which templates S and T had at least two identical non-nil slot values, as well as for the feature modeling an exact match of complex name values. As one might expect, a negative value was learned for the case in which template T was created from an indefinite expression. A positive value was learned for the case in which template T was created from a definite expression and S was (perhaps transitively) the preferred referent according to the coreference module. Interestingly, no value was learned for template S being a possible but non-preferred referent, but a small positive value was learned for it not being on the list at all – presumably this covers cases in which the coreference module fails to identify an existing referent. All the distance features except for *close* and *mid-distance* received negative λ_i values, suggesting that coreference between *close* and *mid-distance* templates was more likely than coreference between templates that were *very close*, *far away*, and *very far away*.

The cross-entropy of the learned model as applied to the training data in each case was about 0.80. Given that the cross-entropy of the uniform distribution and the data is 1 (as there are only two possible values for the random variable, i.e., S and T are coreferent or not), this relatively small reduction suggests that the problem has some amount of difficulty, which is consistent with the notable lack of clear signals of coreference characteristic of the texts in our domain.

Data for the Merging Decision Model Unlike the evidential model, the merging decision model does not always utilize all of the pairwise probabilities between pairs in a coreference set. For instance, in determining the probability of a coreference configuration ((A B C)), it does not consider the probability assigned to the pair A and C except to check that they are compatible. Therefore, the training set for the maximum entropy algorithm was pared down to only contain those pairs that the merger would have considered in deriving the correct coreference configurations. The resulting data had the same coreference sets as the training data for the characteristic of context paired with the result of coreference.

evidential approach, but consisted of characteristics of context for 415 template pairs in the first training set, 405 pairs in the second training set, and 370 pairs in the third training set. The features selected were similar to those in the training of the evidential model.

The cross-entropies of the learned maximum entropy models and the training data were notably better than those for the evidential model, at about 0.70 in each case. This improvement is not particularly surprising. In the evidential case, the fact that all pairs of templates are considered results in a certain amount of “washing out” of the data, due to redundancy in coreference relationships. For instance, coreference between two templates that are far away might be unlikely if there are no corefering expressions between them, but quite likely if there are. When just considering the pairwise feature sets, these two cases are not distinguished, so the resulting probability will be mixed. However, in the merging decision case, pairs that are far away will not be in the data set if there are corefering expressions between them, and thus the probability for coreference at long distances will be diminished. The result is a “cleaner” set of data in which clearer distinctions may be found, as evidenced by the lower cross-entropy achieved.

5.2 Evaluation Results

The cross-entropies of the various approaches as applied to the three sets of test data are shown in Table 2. The number within parentheses indicates the number of times that the coreference set with the highest probability was the correct one. As hoped, both the evidential and merging decision approaches outperformed the uniform and greedy approaches with respect to cross-entropy.¹⁰

Interestingly, and perhaps surprisingly, the evidential approach outperformed the merging decision model, even though in many respects the latter is more natural and elegant. While considering feature sets for all pairs may wash out the training data for the pairwise probability model somewhat, the evidence provided by all pairs appears to more than make up for the difference. Given that a goal of these experiments is to see how well the strategies would perform with a fairly crude, easily computable, and portable set of characteristics of con-

¹⁰The merging decision approach did not do any better than the greedy approach in terms of raw accuracy, and in fact did somewhat worse in the third test. Again, however, the reduction in cross-entropy is important, as the statistics produced by the system will be integrated with other probabilistic factors in the downstream system.

	Test Set 1	Test Set 2	Test Sets 1 and 2
Uniform	2.12 (—)	1.76 (—)	2.01 (—)
Greedy	1.50 (15)	1.30 (20)	1.41 (30)
Merging Decision	1.32 (15)	1.13 (20)	1.27 (27)
Evidential	1.10 (17)	0.89 (21)	1.00 (35)

Table 2: Initial Evaluation Cross-Entropies

text, we are encouraged by the results of these experiments, especially considering the limited amount of training data that was available.

Nonetheless, additional data is necessary to confirm the results of these initial evaluations. Although the consistency of the results between the first two training/test divisions may suggest that the amount of training data is sufficient for the rather coarsely grained feature set used, the size of the test sets are potentially of concern, which motivated our inclusion of the third training/test division. Despite the reduction in training data and corresponding increase in test data, the results of this experiment appear to consistent with the first two.

There are a variety of characteristics of context that one might add to improve the models. For instance, one could add a characteristic indicating when a template is created from a phrase in a subject line or table, as many cases of coreference with subsequent indefinite phrases occur in this circumstance. Other types of information about text type, text structure, and more finely grained distinctions with respect to referential types (e.g., modeling pronouns differently than other definite NPs) would all likely further improve the model, although for some of these additional training data would be required and more domain and genre dependence may result.

While this work was motivated by a need to pass probabilistic output to a downstream data fusion system, these methods can be applied system internally also, to supplant existing algorithms for merging in IE settings that do not allow for probabilistic output. In this scenario, the system simply performs the template merging dictated by the most probable coreference configuration for a given coreference set. However, as noted earlier, the texts in our application are relatively short, and therefore the coreference sets are usually of manageable size. Significantly larger coreference sets can lead to an enormous number of possible coreference configurations. Therefore, to address this task in applications with much longer texts, mechanisms beyond those that were necessary here will be required for intelligently pruning the search space and subsequently smooth-

ing the distributions.

6 Conclusions

Certain applications require that the output of an information extraction system be probabilistic, so that a downstream system can reliably *fuse* the output with possibly contradictory information from other sources. In this paper we considered the problem of assigning a probability distribution to alternative sets of coreference relationships among entity descriptions. We presented the encouraging results of initial experiments with several approaches to estimating such distributions in an application using SRI’s FASTUS information extraction system. We would expect further gains from encoding additional training data and modeling more informative characteristics of context.

Acknowledgments

The author thanks John Bear, Joshua Goodman, and two anonymous reviewers for helpful comments and criticisms, and the SRI Message Handler project team for their contributions to the system in which this work is embedded. This work was supported by the Defense Advanced Research Projects Agency under contract number 4099SCL001 (E-Systems Inc., prime contractor).

References

- Aone, Chinatsu and Scott William Bennett. 1995. Evaluating automated and manual acquisition of anaphora resolution strategies. In *Proceedings of the 33rd Annual Meeting of the Association for Computational Linguistics (ACL-95)*, pages 122–129, Cambridge, MA, June.
- Berger, Adam, Stephen A. Della Pietra, and Vincent J. Della Pietra. 1996. A maximum entropy approach to natural language processing. *Computational Linguistics*, 22(1):39–71.
- Connolly, Dennis, John D. Burger, and David S. Day. 1994. A machine learning approach to

- anaphoric reference. In *Proceedings of the International Conference on New Methods in Language Processing (NeMLaP)*.
- Dagan, Ido and Alon Itai. 1990. Automatic acquisition of constraints for the resolution of anaphora references and syntactic ambiguities. In *Proceedings of the 13th International Conference on Computational Linguistics (COLING-90)*, pages 330–332.
- Dagan, Ido, John Justenson, Shalom Lappin, Herbert Leass, and Amnon Ribak. 1995. Syntax and lexical statistics in anaphora resolution. *Applied Artificial Intelligence*, 9(6):633–644, Nov/Dec.
- Dempster, Arthur P. 1968. A generalization of Bayesian inference. *Journal of the Royal Statistical Society*, 30:205–247.
- Hobbs, Jerry R., Douglas E. Appelt, John Bear, David Israel, Megumi Kameyama, Mark Stickel, and Mabry Tyson. 1996. FASTUS: A cascaded finite-state transducer for extracting information from natural-language text. In *Finite State Devices for Natural Language Processing*. MIT Press, Cambridge, MA.
- Kennedy, Christopher and Branimir Boguraev. 1996a. Anaphora for everyone: Pronominal anaphora resolution without a parser. In *Proceedings of the 16th International Conference on Computational Linguistics (COLING-96)*.
- Kennedy, Christopher and Branimir Boguraev. 1996b. Anaphora in a wider context: Tracking discourse referents. In *Proceedings of the 12th European Conference on Artificial Intelligence (ECAI-96)*.
- Lappin, Shalom and Herbert Leass. 1994. An algorithm for pronominal anaphora resolution. *Computational Linguistics*, 20(4):535–561.